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Hierarchical Bayesian model of population structure reveals convergent adaptation to high altitude in human populations

Running head: Convergent adaptation to high altitude in humans

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Abstract

Detecting genes involved in local adaptation is challenging and of fundamental importance in evolutionary, quantitative, and medical genetics. To this aim, a standard strategy is to perform genome scans in populations of different origins and environments, looking for genomic regions of high differentiation. Because shared population history or population sub-structure may lead to an excess of false positives, analyses are often done on multiple pairs of populations, which leads to i) a global loss of power as compared to a global analysis, and ii) the need for multiple tests corrections. In order to alleviate these problems, we introduce a new hierarchical Bayesian method to detect markers under selection that can deal with complex demographic histories, where sampled populations share part of their history. Simulations show that our approach is both more powerful and less prone to false positive loci than approaches based on separate analyses of pairs of populations or those ignoring existing complex structures. In addition, our method can identify selection occurring at different levels (i.e. population or region-specific adaptation), as well as convergent selection in different regions. We apply our approach to the analysis of a large SNP dataset from low- and high-altitude human populations from America and Asia. The simultaneous analysis of these two geographic areas allows us to identify several new candidate genome regions for altitudinal selection, and we show that convergent evolution among continents has been quite common. In addition to identifying several genes and biological processes involved in high altitude adaptation, we identify two specific biological pathways that could have evolved in both continents to counter toxic effects induced by hypoxia.

Author Summary

We present a new Bayesian genome scan method to detect markers under natural selection from multiple populations living in different environments. Our method can deal with complex demographic histories, where populations sampled in a given geographic region share part of their history, and can identify selection occurring at different levels (i.e. population or region-specific adaptation), as well as convergent adaptation in different regions. Simulations show that our approach is both more powerful and less prone to false positives than approaches based on separate analyses of pairs of populations or those ignoring existing complex structures. We apply this approach to a large genomic dataset from low- and high-altitude human populations from America and Asia. We identify several new candidate genome regions for altitudinal selection, and we show that convergent adaptation among continents is much more common than previously thought. In addition to identifying several genes and biological processes involved in high altitude adaptation, we identify two specific biological pathways that could have evolved in both continents to counter toxic effects induced by hypoxia.

Introduction

Distinguishing between neutral and selected molecular variation has been a long-standing interest of population geneticists. This interest was fostered by the publication of Kimura's seminal paper [1] on the neutral theory of molecular evolution. Although the controversy rests mainly on the relative importance of genetic drift and selection as explanatory processes for the observed biodiversity patterns, another important question concerns the prevalent form of natural selection. Kimura [1] argued that the main selective mechanism was negative selection against deleterious mutations. However, an alternative point of view

emphasizes the prevalence of positive selection, the mechanism that can lead to local adaptation and eventually to speciation [2,3].

A powerful approach to uncover positive selection is the study of mechanisms underlying convergent evolution. When different populations or evolutionary lineages are exposed to similar environments, positive selection should indeed lead to similar phenotypic features. Convergent evolution can be achieved through similar genetic changes (sometimes called “parallel evolution”) at different levels: the same mutation appearing independently in different populations, the same existing mutation being recruited by selection in different populations, or the involvement of different mutations in the same genes or the same biological pathways in separate populations [4]. However, existing statistical genetic methods are not well adapted to the study of convergent evolution when data sets consists in multiple contrasts of populations living in different environments [5]. The current strategy is to carry out independent genome scans in each geographic region and to look for overlaps between loci or pathways that are identified as outliers in different regions [6]. Furthermore, studies are often split into a series of pairwise analyses that consider sets of populations inhabiting different environments. Whereas this strategy has the advantage of not requiring the modeling of complex demographic histories [7,8], it often ignores the correlation in gene frequencies between geographical regions when correcting for multiple tests [9]. As a consequence, current approaches are restricted to the comparison of lists of candidate SNPs or genomic regions obtained from multiple pairwise comparisons. This sub-optimal approach may also result in a global loss of power as compared to a single global analysis and thus to a possible underestimation of the genome-wide prevalence of convergent adaptation.

One particularly important example where this type of problems arises is in the study of local adaptation to high altitude in humans. Human populations living at high altitude need to cope with one of the most stressful environment in the world, to which they are likely to have developed specific adaptations. The harsh conditions associated with high altitude do not only include low oxygen partial pressure, referred to as high-altitude hypoxia, but also encompass other factors like low temperatures, arid climate, high solar radiation and low soil quality. While some of these stresses can be buffered by behavioral and cultural adjustments, important physiological changes have been identified in populations living at high altitude (see below). Recently, genomic advances have unveiled the first genetic bases of these physiological changes in Tibetan, Andean and Ethiopian populations [10-19]. The study of convergent or independent adaptation to altitude is of primary interest [11,19,20], but this problem has been superficially addressed so far, as most studies focused on populations from a single geographical region [10,13,14,16-19].

Several candidate genes for adaptation to altitude have nevertheless been clearly identified [21,22], the most prominent ones being involved in the hypoxia inducible factor (HIF) pathway, which plays a major role in response to hypoxia [23]. In Andeans, *VEGF* (vascular endothelial growth factor), *PRKAA1* (protein kinase, AMP-activated, alpha 1 catalytic subunit) and *NOS2A* (nitric oxide synthase 2) are the best-supported candidates, as well as *EGLN1* (egl-9 family hypoxia-inducible factor 1), a down regulator of some HIF targets [12,24]. In Tibetans [10,11,13,14,16,25], the HIF pathway gene *EPAS1* (endothelial PAS domain protein 1) and *EGLN1* have been repeatedly identified [22]. Recently, three similar studies focused on Ethiopian highlanders [17-19] suggested the involvement of HIF genes other than those identified in Tibetans and Andeans, with *BHLHE41*, *THRB*, *RORA* and *ARNT2* being the most prominent candidates.

However, there is little overlap in the list of significant genes in these three regions [18,19], with perhaps the exception of alcohol dehydrogenase genes identified in two out of the three analyses. Another exception is *EGLN1*: a comparative analysis of Tibetan and Andean populations [12] concluded that “the Tibetan and Andean patterns of genetic adaptation are largely distinct from one another”, identifying a single gene (*EGLN1*) under convergent evolution, but with both populations exhibiting a distinct dominant haplotype around this gene. This limited convergence does not contradict available physiological data, as Tibetans exhibit some phenotypic traits that are not found in Andeans [26]. For example, Tibetan populations show lower hemoglobin concentration and oxygen saturation than Andean populations at the same altitude [27]. Andeans and Tibetans also differ in their hypoxic ventilatory response, birth weight and pulmonary hypertension [28]. Finally, *EGLN1* has also been identified as a candidate gene in Kubachians, a high altitude (~2000 m a. s. l.) Daghestani population from the Caucasus [15], as well as in Indians [29].

Nevertheless, it is still possible that the small number of genes under convergent evolution is due to a lack of power of genome scan methods done on separate pairs of populations. In order to overcome these difficulties, we introduce here a new Bayesian genome scan method that (i) extends the F-model [30,31] to the case of a hierarchically subdivided population consisting of several migrant pools, and (ii) explicitly includes a convergent selection model. We apply this new approach to find genes, genomic regions, and biological pathways that have responded to convergent selection in the Himalayas and in the Andes.

Material and methods

Hierarchical Bayesian model

One of the most widely used statistics for the comparison of allele frequencies among populations is F_{ST} [31-34], and all studies cited in the introduction used it to compare low- and high altitude populations within a given geographical region (Tibet, the Andes or Ethiopia). Several methods have been proposed to detect loci under selection from F_{ST} , and one of the most powerful approach is based on the F-model [35]. However, this approach assumes a simple island model where populations exchange individuals through a unique pool of migrants. This assumption is strongly violated when dealing with replicated pairs of populations across different regions, which can lead to a high rate of false positives [34].

In order to relax the rather unrealistic assumption of a unique and common pool of migrants for all sampled populations, we extended the genome scan method first introduced by Beaumont and Balding [30] and later improved by Foll and Gaggiotti [31]. More precisely, we posit that our data come from G groups (migrant pools or geographic regions), each group g containing J_g populations. We then describe the genetic structure by a F-model that assumes that allele frequencies at locus i in population j from group g ,

$\mathbf{p}_{ijg} = \{p_{ijg1}, p_{ijg2}, \dots, p_{ijgK_i}\}$ (where K_i is the number of distinct alleles at locus i), follow a

Dirichlet distribution parameterized with group-specific allele frequencies

$\mathbf{p}_{ig} = \{p_{ig1}, p_{ig2}, \dots, p_{igK_i}\}$ and with F_{SC}^{ijg} coefficients measuring the extent of genetic

differentiation of population j relative to group g at locus i . Similarly, at a higher group

level, we consider an additional F-model where allele frequencies \mathbf{p}_{ig} follow a Dirichlet

distribution parameterized with meta-population allele frequencies $\mathbf{p}_i = \{p_{i1}, p_{i2}, \dots, p_{iK_i}\}$

and with F_{CT}^{ig} coefficients measuring the extent of genetic differentiation of group g relative

to the meta-population as a whole at locus i . Figure S1 shows the hierarchical structure of

our model in the case of three groups ($G = 3$) and four populations per group ($J_1 = J_2 = J_3 = 4$) and Figure S2 shows the corresponding non-hierarchical F-model for the

same number of populations. All the parameters of the hierarchical model can be estimated

by further assuming that alleles in each population j are sampled from a Multinomial

distribution [36]. These assumptions lead to an expression for the probability of observed

allele counts $\mathbf{a}_{ijg} = \{a_{ijg1}, a_{ijg2}, \dots, a_{ijgK_i}\}$:

$$\Pr(\mathbf{a}_{ijg} | \mathbf{p}_{ijg}, \mathbf{p}_{ig}, \mathbf{p}_i, \theta_{ijg}, \phi_{ig}) = \Pr(\mathbf{a}_{ijg} | \mathbf{p}_{ijg}) \Pr(\mathbf{p}_{ijg} | \mathbf{p}_{ig}, \theta_{ijg}) \Pr(\mathbf{p}_{ig} | \mathbf{p}_i, \phi_{ig})$$

where $\Pr(\mathbf{a}_{ijg} | \mathbf{p}_{ijg})$ is the Multinomial likelihood, $\Pr(\mathbf{p}_{ijg} | \mathbf{p}_{ig}, \theta_{ijg})$ and $\Pr(\mathbf{p}_{ig} | \mathbf{p}_i, \phi_{ig})$ are

Dirichlet prior distributions, $\theta_{ijg} = 1 / F_{SC}^{ijg} - 1$, and $\phi_{ig} = 1 / F_{CT}^{ig} - 1$. This expression can be

simplified by integrating over \mathbf{p}_{ijg} so as to obtain:

$$\Pr(\mathbf{a}_{ijg} | \mathbf{p}_{ig}, \mathbf{p}_i, \theta_{ijg}, \phi_{ig}) = \Pr(\mathbf{a}_{ijg} | \mathbf{p}_{ig}, \theta_{ijg}) \Pr(\mathbf{p}_{ig} | \mathbf{p}_i, \phi_{ig})$$

where $\Pr(\mathbf{a}_{ijg} | \mathbf{p}_{ig}, \theta_{ijg})$ is the Multinomial-Dirichlet distribution [35]. The likelihood is

obtained by multiplying across loci, regions and population

$$L(\mathbf{p}_{ig}, \mathbf{p}_i, \theta_{ijg}, \phi_{ig}) = \prod_{i=1}^L \prod_{g=1}^G \prod_{j=1}^{J_g} \Pr(\mathbf{a}_{ijg} | \mathbf{p}_{ig}, \mathbf{p}_i, \theta_{ijg}, \phi_{ig})$$

Using this model, we incorporate potential deviation from the genome wide F-statistics at each locus as in Beaumont and Balding [30]. The genetic differentiation within each group g is:

$$\log\left(\frac{F_{SC}^{ijg}}{1 - F_{SC}^{ijg}}\right) = \alpha_{ig} + \beta_{jg} \quad (1)$$

where α_{ig} is a locus-specific component of F_{SC}^{ijg} shared by all populations in group g , and β_{jg} is a population-specific component shared by all loci. Similarly, we decompose the genetic differentiation at the group level under a logistic model as:

$$\log\left(\frac{F_{CT}^{ig}}{1 - F_{CT}^{ig}}\right) = A_i + B_g \quad (2)$$

where A_i is a locus-specific component of F_{CT}^{ig} shared by all groups in the meta-population, and B_g is a group-specific component shared by all loci.

By doing this, our model also eliminates the ambiguity of having a single α_i parameter for more than two populations, since we now have (i) different selection parameters in each geographic region (α_{ig} are group specific) and (ii) separate parameter sensitive to adaptation among regions at the larger scale (A_i). We use the likelihood function and the logistic decomposition to derive the full Bayesian posterior:

$$\Pr(\mathbf{p}_{ig}, \mathbf{p}_i, A_{ig}, B_{jg}, \alpha_i, \beta_g | \mathbf{A}) \propto L(\mathbf{p}_{ig}, \mathbf{p}_i, \theta_{ijg}, \phi_{ig}) \Pr(\mathbf{p}_{ig} | \mathbf{p}_i, \alpha_i, \beta_g) \Pr(\mathbf{p}_i) \Pr(A_{ig}) \Pr(B_{jg}) \Pr(\alpha_i) \Pr(\beta_g)$$

where the prior for \mathbf{p}_i is a non-informative Dirichlet distribution, the priors for α_{ig} and A_i are Gaussian with mean 0 and variance 1, and the priors for β_{jg} and B_g are Gaussian with mean -1

and variance 1. Note that priors on β_{jg} and B_g have practically no influence on the posteriors as these parameter use the huge information coming from all loci.

Parameter estimation

We extend the Reversible Jump Markov Chain Monte Carlo (RJ-MCMC) approach proposed by Foll and Gaggiotti [31] to identify selection both within groups and among groups. For each locus and in each group separately, we consider a neutral model where $\alpha_{ig} = 0$, and a model with selection where the locus-specific effect $\alpha_{ig} \neq 0$. Similarly, we consider two models at the group level for each locus where $A_i = 0$ for the neutral model, and $A_i \neq 0$ for the model with selection. In order to tailor our approach to study convergent adaptation, we also consider the case where different groups share a unique locus-specific component α_i (see Figure 1 for an example of such a model with two groups of two populations). At each iteration of the MCMC algorithm, we update A_i and α_{ig} in a randomly chosen group g for all loci. As described in [31], we propose to remove α_{ig} from the model if it is currently present, or to add if it is not, and we do the same for A_i . We also add a new specific Reversible Jump proposal for convergent adaptation: if all groups but one are currently in the selection model ($\alpha_{ig} \neq 0$ for all g but one), we propose with a probability 0.5 to move to the convergent evolution model (where we replace all α_{ig} by a single selection parameter α_i shared by all groups), and with a probability 0.5 we perform a standard jump as described above.

Genomic data set

In order to improve our understanding of the genetic bases of adaptation to altitude, we have applied our new hierarchical Bayesian method to the dataset from Bigham *et al.* [12]. This data set includes four populations genotyped for 906,600 SNPs using the Affymetrix Genome-Wide Human SNP Array 6.0 platform (<http://www.affymetrix.com>). It includes two populations living at high altitude in the Andes (49 individuals) and in Tibet (49 individuals), as well as two lowland related populations from Central-America (39 Mesoamericans) and East Asia (90 individuals from the international HapMap [37] project). Thus, we compared four alternative models for each locus at the population level: 1) a neutral model ($\alpha_{i1} = \alpha_{i2} = 0$), 2) a model with selection acting only in Tibetans ($\alpha_{i2} = 0$), 3) a model with selection acting only in Andeans ($\alpha_{i1} = 0$), and 4) a convergent adaptation model with selection acting similarly in both Tibetans and Andeans ($\alpha_{i1} = \alpha_{i2} = \alpha_i$). We estimate the posterior probability that a locus is under selection by summing up the posterior probabilities of the three non-neutral models (2, 3 and 4) and we control for False Discovery Rate (FDR) by calculating associated q -values [38-40]. We do not pay any particular attention to the A_i parameter here, as it can be interpreted as a potential adaptation at the continental level in Asians and Native Americans, which is not directly relevant in the context of adaptation to high altitude (but see Discussion).

We excluded SNPs with a global minor allele frequency below 5% to avoid potential biases due to uninformative polymorphisms [41]. This left us with 632,344 SNPs that were analyzed using the hierarchical F-model described above. We identified genomic regions potentially involved in high altitude adaptation by using a sliding windows approach. We considered windows of 500 kb, with a shifting increment of 25 kb at each step. The average number of

SNPs per window over the whole genome was 121.4 (sd=44.6), after discarding any window containing less than 50 SNPs. We considered a window as a candidate target for selection if the 5% quantile of the q -values in the window was lower than 0.01, and we merged overlapping significant windows into larger regions.

Detecting polygenic convergent adaptation

We first used SNPs identified as being under convergent adaptation to perform classical enrichment tests for pathways using Panther [42] and Gene Ontology (GO) [43] using String 9.1 [44]. More specifically, we extracted the list of 302 genes within 10 kb of all SNPs assigned to the convergent adaptation model and showing a q -value below 10%, to serve as input for these tests.

These two approaches have limited power to detect selection acting on polygenic traits, for which adaptive events may have arisen from standing variation rather than from new mutations [3,25]. In order to detect polygenic adaptation, we used a new gene set enrichment approach [45], which tests if the distribution of a statistic computed across genes of a given gene set is significantly different from the rest of the genome. As opposed to the classical enrichment tests, this method does not rely on an arbitrary threshold to define the top outliers and it uses all genes that include at least one tested SNP. In short, we tested more than 1,300 gene sets listed in the Biosystems database [46] for a significant shift in their distribution of selection scores relative to the baseline genome-wide distribution. In our case, the selection score of each SNP is its associated q -value of convergent selection. As previously done [45], we calculated the sum of gene scores for each gene set and compared it to a null distribution of random sets ($N=500,000$) to infer its significance. In order to avoid any redundancy between gene sets, we iteratively removed genes belonging to the most

significant gene sets from the less significant gene sets before testing them again in a process called “pruning”. This process leads to a list of gene sets whose significance is obtained by testing completely non-overlapping lists of genes. See the Supplemental Material Text S1 for a more detailed description of the method.

Simulation

In order to evaluate the performance of our hierarchical method, we simulated data with features similar to the genomic data set analyzed here under our hierarchical F-model. Our simulated scenario thus includes two groups of two populations made of 50 diploids each, with $F_{SC} = 0.02$ for all four populations and $F_{CT} = 0.08$ for both groups. In each group, a fraction of loci are under selective pressure in one of the two populations only. We simulated a total of 100,000 independent SNPs among which (i) 2,500 are under weak convergent evolution with $\alpha_i = 3$, (ii) 2,500 are under stronger convergent evolution with $\alpha_i = 5$, (iii) 2,500 are under weak selection in the first group with $\alpha_{i1} = 3$ and neutral ($\alpha_{i2} = 0$) in the second group, (iv) 2,500 are under stronger selection in the first group with $\alpha_{i1} = 5$ and neutral ($\alpha_{i2} = 0$) in the second group, and (v) 90,000 remaining SNPs that are completely neutral ($\alpha_{i1} = \alpha_{i2} = 0$). As in the real data, we conditioned the SNPs to have a global minor allele frequency above 5%. We analyzed this simulated dataset using three different approaches: (i) the new hierarchical F-model introduced above, (ii) two separate pairwise analyses (one for each group) containing two populations using the original F-model implemented in BayeScan [31], (iii) a single analysis containing the four populations using the original F-model implemented in BayeScan [31] ignoring the hierarchical structure of the populations. In our hierarchical model, the best selection model for each SNP was identified as

described above using a $q\text{-value} < 0.01$. When analyzing data in separate pairs of populations, we considered a SNP to be under convergent adaptation when it had a $q\text{-value} < 0.01$ in the two regions.

Results

Patterns of selection at the SNP level

Using our hierarchical Bayesian analysis, we identified 1,159 SNPs potentially under selection at the 1% FDR level ($q\text{-value} < 0.01$). For each SNP, we identified the model of selection (selection only in Asia, selection only in South America, or convergent selection; see methods) with the highest posterior probability. With this procedure, 362 SNPs (31%) were found under convergent adaptation, whereas 611 SNPs (53%) were found under selection only in Asia, and 186 SNPs (16%) only under selection in South America. These results suggest that convergent adaptation is relatively common, even at the SNP level, at odds with results from previous analyzes [5,24,32], but consistent with results of a recent literature meta-analysis over several species [5].

In order to evaluate the additional power gained with the simultaneous analysis of the four populations, we performed separate analyses in the two continents using the non-hierarchical F-model [31]. These two pairwise comparisons identified 160 SNPs under selection in the Andes, and 940 in Tibet. The overlap in significant SNPs between these two separate analyzes and that under the hierarchical model is shown in Figure 2A. Interestingly, only 6 SNPs are found under selection in both regions when performing separate analyses in Asians and Amerindians. This very limited overlap persists even if we strongly relax the FDR in both pairwise analyzes: at the 10% FDR level only 13 SNPs are found under selection in

both continents. These results are consistent with those of Bigham *et al.* [12], who analyzed both continents separately with a different statistical method based on F_{ST} . This suggests that the use of intersecting sets of SNPs found significant in separate analyses is not a suitable strategy to study the genome-wide importance of convergent adaptation. Interestingly, 15% of the SNPs (162 SNPs, see Figure 2A) identified as under selection by our new method are not identified by any separate analyses, showing the net gain of power of our method for detecting genes under selection, as confirmed by our simulation study below.

We examined in more detail the 362 SNPs identified as under convergent adaptation. The overlap of these SNPs (the yellow circle) with those identified by the two separate analyses is shown in Figure 2B. As expected, the 6 SNPs identified under selection in both regions by the two separate analyses are part of the convergent adaptation set. However, we note that 272 of the SNPs in the convergent adaptation set (75%) are identified as being under selection in only one of the two regions by the separate analyses. This suggests that although natural selection may be operating similarly in both regions, limited sample size may prevent its detection in one of the two continents.

Genomic regions under selection

Using a sliding window approach, we find 25 candidate regions with length ranging from 500 kb to 2 Mb (Figure 3 and Table S1). Among these, 18 regions contain at least one significant SNP assigned to the convergent adaptation model, and 11 regions contained at least one 500 kb-window where the convergent adaptation model was the most frequently assigned selection model among significant SNPs (Figure 3). Contrastingly, Bigham *et al.* [12] identified 14 and 37 candidate 1 Mb regions for selection in Tibetans and Andeans,

respectively, but none of these 1 Mb regions were shared between Asians and Amerindians. Moreover, only two of the regions previously found under positive selection in South America and only four in Asia overlap with our 25 significant regions.

As noted above, the only gene showing signs of convergent evolution found by Bigham *et al.* [12] is *EGLN1*, which has also been identified in several other studies (see Table 1 in [22] for a review). *EGLN1* is also present in one of our 25 regions where three out of eight significant SNPs are assigned to the convergent adaptation model. We note that the significant SNPs in this region are not found in *EGLN1* directly but in two genes surrounding it (*TRIM67* and *DISC1*), as reported earlier [14,47]. The HIF pathway gene *EPAS1*, which is the top candidate in many studies [22], is also present in one of our 25 regions, where 28 of the 80 significant SNPs are assigned to the convergent adaptation model.

Polygenic convergent adaptation

We identified three pathways significantly enriched for genes involved in convergent adaptation using Panther [42] after Bonferroni correction at the 5% level. These are the “metabotropic glutamate receptor group I” pathway, the “muscarinic acetylcholine receptor 1 and 3” signaling pathway, and the “epidermal growth factor receptor” (EGFR) signaling pathway. Using the String 9.1 database [44], two GO terms were significantly enriched for these genes when controlling for a 5% FDR: “ethanol oxidation” (GO:0006069) and “positive regulation of transmission of nerve impulse” (GO:0051971). Using a new and more powerful gene set enrichment approach [45], we first identified 25 gene sets with an associated *q*-value below 5% (Table S2). An enrichment map showing these sets and their overlap is presented in Figure 4. There are two big clusters of overlapping gene sets, one related to Fatty Acid Oxidation with “Fatty Acid Omega Oxidation” as the most significant set and

another immune system related cluster with "*Interferon gamma signaling*" as the most significant gene set. After pruning, only these two above-mentioned gene sets are left with a q -value below 5%. It is worth noting that the "*Fatty Acid Omega Oxidation*" pathway, which is the most significant gene set ($q\text{-value} < 10^{-6}$), contains many top scoring genes for convergent selection, including several alcohol and aldehyde dehydrogenases, as listed in Table S3. Interestingly, the GO term "*ethanol oxidation*" is no longer significant after excluding the genes involved in the "*Fatty Acid Omega Oxidation*" pathway.

Power of the hierarchical F-model

Our simulations show a net increase in power to detect selection using the global hierarchical approach as compared to using two separate pairwise analyses (Table 1 and Figure 5 and 6). For the 2,500 SNPs simulated under the weak convergent selection model ($\alpha_i = 3$), the hierarchical model detects 6.5 times more SNPs than the two separate analyses (306 vs. 47). The power greatly increases when selection is stronger, and among the 2,500 SNPs simulated with $\alpha_i = 5$, 1,515 are correctly classified using our hierarchical model, as compared to only 643 using separate analyses. Similarly to what we found with the real altitude data, the two separate analyses often wrongly classify the convergent SNPs correctly identified as such by our hierarchical method as being under selection in only one of the two groups, but sometimes also as completely neutral (64 such SNPs when $\alpha_i = 3$ and 76 when $\alpha_i = 5$, see Figure 5B and D). We note that the hierarchical model is also more powerful at detecting selected loci regardless of whether or not the SNPs are correctly assigned to the convergent evolution set. Indeed, the new method identifies 2,626 SNPs as being under any model of selection (*i. e.* convergent evolution or in only one of the two regions) among the 5,000 simulated under convergent selection, whereas the separate analysis detects only

2,475 SNPs. When selection is present only in one of the two groups ($\alpha_{i1} = 3$ or 5 and $\alpha_{i2} = 0$), the power of the hierarchical model is comparable with the separate analysis in the corresponding group, implying that there is no penalty in using the hierarchical model even in presence of group specific selection. A few of the group-specific selected SNPs are wrongly classified in the convergent adaptation model with a false positive rate of 1.7% (84 SNPs out of 5,000). Overall, the false discovery rate is well calibrated using our q -value threshold of 0.01 in both cases, with 29 false positives out of 4,141 significant SNPs (FDR=0.70%) for our hierarchical model, and 30 false positives out of 3,984 significant SNPs (FDR=0.75%) for the two separate analyses. Finally, when the four populations are analyzed together without accounting for the hierarchical structure, a large number of false positives appears (Table 1 and Figure 6C) in keeping with previous studies [34]. Under this island model, 1,139 neutral SNPs are indeed identified as being under selection among the 90,000 simulated neutral SNPs (vs. 29 and 30 using the hierarchical method or two separate analyses, respectively). The non-hierarchical approach does not allow one to distinguish different models of selection, but among the 10,000 SNPs simulated under different types of selection, only 2,598 are significant. This shows that the non-hierarchical analysis leads to both a reduced power, and a very large false discovery rate (FDR=30.4%) in presence of a hierarchical genetic structure.

Discussion

Convergent adaptation to high altitude in Asia and America is not rare

Our new hierarchical F-model reveals that convergent adaptation to high altitude is more frequent than previously described in Tibetans and Andeans. Indeed, 31% (362/1159) of all SNPs found to be potentially under selection at a FDR of 1% can be considered as under

convergent adaptation in Asia and America. This is in sharp contrast with a previous analysis of the same data set where only a single gene was found to be responding to altitudinal selection in both Asians and Amerindians [12]. Our new model confirms the selection of *EGLN1* in both Tibetans and Andeans. We also show that some genes already known to be involved in adaptation to high altitude in Tibetans, like the *EPAS* gene, may also have the same function in Andeans. Finally, we identified genomic regions, pathways, and GO terms potentially linked to convergent adaptation to high altitude in Tibetans and Andeans that have not been previously reported. Our approach seems thus more powerful than previous pairwise analyses, which is confirmed by our simulation study. It suggests that datasets analyzed by previous studies that tried to uncover convergent adaptation by confronting lists of significant SNPs in separate pairwise analyses [48-52] would benefit from being reanalyzed with our method.

Polygenic and convergent adaptation in the omega oxidation pathway

Our top significant GO term is linked to alcohol metabolism, in keeping with a recent study of a high altitude population in Ethiopia [18,19]. Indeed, one of the 25 regions identified in the present study includes several alcohol dehydrogenase (ADH) genes (ADH1A-B-C and ADH4-5-6-7) located in a 370 kb segment of chromosome 4 (Figure 3), and another significant segment of 2 Mb portion of chromosome 12 includes the acetaldehyde dehydrogenase ALDH2 gene. Some evidence of positive selection in the ADH1B and ALDH2 genes had been reported in East-Asian populations. This signal was linked to the expansion of rice domestication and fermentation [53-55], because plants show a higher rate of ethanol fermentation under low oxygen stress conditions [56], and a similar adaptation to alcoholic fermentation has been previously reported in *Drosophila* [57].

Interestingly, our gene set enrichment analysis provides additional insights for a potential evolutionary adaptation of this group of genes, since they all belong to the most significant identified pathway, namely "*Fatty Acid Omega Oxidation*" (Table S3). Omega oxidation is an alternative to the beta-oxidation pathway by which fatty acids are transformed into succinate that can enter the citric acid cycle and produce energy in the mitochondrion. This way to degrade fatty acids into sugar is usually a minor metabolic pathway, which becomes more important when beta-oxidation is defective, like in the case of fatty acid oxidation disorder in humans [18,19,58], or in case of hypoxia [59]. Note that glucose oxidation and glycolysis have been shown to be a more efficient source of energy production than fatty acid oxidation in Tibetans, leading to an increased lactase and free fatty acid concentration [60]. It is thus unclear if omega oxidation is a more efficient alternative to beta oxidation at high altitude, or if it would rather contribute to the degradation of fatty acids accumulating when beta oxidation is defective. The detoxifying role of this pathway is supported by the fact that it is usually mainly active in the liver and in the kidney [58]. The fact that Ethiopians also show signals of adaptations in ADH and ALDH genes [19] suggests that convergent adaptation in the omega oxidation pathway could have occurred on three different continents in humans.

Response to hypoxia-induced neuronal damage

Hypoxia leads to neuronal damage through over-stimulation of glutamate (an amino acid acting as an excitatory neurotransmitter) receptors [61]. Two out of our three significant pathways found with Panther ("*metabotropic glutamate receptor group I*" and "*muscarinic acetylcholine receptor 1 and 3*") for convergent adaptation are involved with neurotransmitter receptors. The metabotropic glutamate receptor group I increases N-

methyl-D-aspartate (*NMDA*) receptor activity, and this excitotoxicity is a major mechanism of neuronal damage and apoptosis [62]. Consistently, the only significant GO term after excluding the genes involved in omega oxidation is also related to neurotransmission (*“positive regulation of transmission of nerve impulse”*) and contains two significant glutamate receptors genes (*GRIK2* and *GRIN2B*) as well as *IL6*. The *GRIN2B* gene encodes one of the four *NMDA NR2* subunits (*NMDAR2B*) that acts as the agonist binding site for glutamate and promotes excitotoxic neuronal apoptosis [63].

One of our top candidate regions for convergent adaptation is located on chromosome 7 and includes 19 significant SNPs assigned to the convergent adaptation model, which are spread in a 100 kb region around *IL6* (Figure 3). This gene encodes interleukin-6 (IL-6), an important cytokine involved in several biological processes. Interestingly it has been shown that IL-6 plasma levels increases significantly when sea-level resident individuals are exposed to high altitude (4300 m) [64]. IL-6 has been shown to have a neuroprotective effect against glutamate- or NMDA-induced excitotoxicity [65]. Additionally, the *“Interferon gamma signaling”* pathway we find using the gene set enrichment approach has been shown to play an important role in neurons and central nervous system reparation after injury [66]. Here again, a consistent result has been identified in Ethiopian highlanders, namely the *“metabotropic glutamate receptor group III”* pathway [17]. Together, these results suggest a genetic adaptive response to neuronal excitotoxicity induced by high altitude hypoxia in humans.

Versatility of the hierarchical Bayesian model to uncover selection

Our statistical model is very flexible and can cope with a variety of sampling strategies to identify adaptation. For example, Pagani *et al.* [15] used a different sampling scheme to

uncover high altitude adaptation genes in North-Caucasian highlanders. They sampled Daghestani from three closely related populations (Avars, Kubachians, and Laks) living at high altitude that they compared with two lowland European populations. Here again, our strategy would allow the incorporation of these five populations into a single analysis. A first group would correspond to the Daghestan region, containing the three populations and a second group containing the two lowland populations. However, in that case, it is the decomposition of F_{CT} in equation 2 that would allow the identification of loci overly differentiated between Daghestani (“group 1”) and European (“group 2”) populations.

Our approach could also be very useful in the context of Genome Wide Association Studies (GWAS) meta-analysis. For example, Scherag *et al.* [67] combined two GWAS on French and German samples to identify SNPs associated with extreme obesity in children. These two data sets could be combined and a single analysis could be performed under our hierarchical framework, explicitly taking into account the population structure. Our two “groups” in Figure 1 would correspond respectively to French and German individuals. In each group the two “populations” would correspond respectively to cases (obese children) and controls (children with normal weight). Like in the present study, the decomposition of F_{SC} and the use of a convergent evolution model would allow the identification of loci associated with obesity in both populations. Additionally, a potential hidden genetic structure between cases and controls and any shared ancestry between French and Germans would be dealt with by the β_{jg} and B_g coefficients in equations 1 and 2, respectively.

Conclusion

We have introduced here a flexible hierarchical Bayesian model that can deal with complex population structure, and which allows the simultaneous analysis of populations living in

different environments in several distinct geographic regions. Our model can be used to specifically test for convergent adaptation, and this approach is shown to be more powerful than previous methods that analyze pairs of populations separately. The application of our method to the detection of loci and pathways under selection reveals that many genes are under convergent selection in the American and Tibetan highlanders and suggests that two specific pathways could have evolved to counter the toxic effects of hypoxia. Future improvements will be directed toward the modeling of genetic linkage between adjacent SNPs, for example using hidden Markov models [68], and the inclusion of shared selective parameters between SNPs located in the same gene or genetic pathways.

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Figure legends

Figure 1. Hierarchical F-model for the high altitude data analyzed.

Directed acyclic graph describing the Bayesian formulation of the hierarchical F-model at a given locus i . Square nodes represent data and circles represent model parameters to be estimated. Dashed circles represent population allele frequencies, which are analytically integrated using a Dirichlet-multinomial distribution (see method description). Lines between the nodes represent direct stochastic relationships within the model. With the exception of Figure 4, we use the same color codes in all Figures, with blue for Asia, red for America, and yellow for convergent adaptation.

Figure 2. Overlap of candidate SNPs under selection in Asia and in America.

Venn diagrams showing the overlap of SNPs potentially under selection in Asia and in America at a 1% FDR. A: Overlap between all SNPs found under any type of selection using our hierarchical model introduced here (green) with those found in separate analyses performed in Asia (blue) and in America (red). B: Overlap between SNPs found under convergent selection using our hierarchical model (yellow) with those found in separate analyses performed in Asia (blue) and in America (red).

Figure 3. Manhattan plot of q -values for loci potentially under altitudinal selection in Asian and Amerindian populations.

Each dot represents the 5% quantile of the SNPs q -values in a 500 kb window. Windows are shifted by increment of 25 kb and considered as a candidate target for selection if the 5% quantile is lower than 0.01 (horizontal dashed line). Overlapping significant windows are merged into 25 larger regions (indicated by grey vertical bars, see Table S2). Significant windows are colored in yellow when they contain at least one significant SNP for convergent adaptation. Otherwise they

are colored according to the most represented model of selection identified among the SNPs they contain: blue for selection only in Asia and red for selection only in America. We also report the names of genes discussed in the text.

Figure 4. Gene sets enriched for signals of convergent adaptation.

The 25 nodes represent gene sets with $q\text{-value} < 0.05$. The size of a node is proportional to the number of genes in a gene set. The node color scale represents gene set p-values. Edges represent mutual overlap: nodes are connected if one of the sets has at least 33% of its genes in common with the other gene set. The widths of the edges scale with the similarity between nodes.

Figure 5. Overlap of significant SNPs for simulated convergent evolution.

Venn diagrams showing the overlap of SNPs simulated under a convergent evolution model and identified under selection at a 1% FDR. A and C: Overlap between SNPs found under any type of selection using our hierarchical model introduced here (green) with those found in separate analyses performed in group 1 (blue) and in group 2 (red). B and D: Overlap between SNPs found under convergent using our hierarchical model (yellow) with those found in separate analyses performed in group 1 (blue) and in group 2 (red). In A and B, 2,500 SNPs are simulated under weak convergent selection ($\alpha_i = 3$), while in C and D 2,500 SNPs are simulated under stronger convergent selection ($\alpha_i = 5$).

Figure 6. Power to detect loci under selection as a function of their effect on population differentiation.

For simulated SNPs, we plot the best selection model inferred (A) under our new hierarchical F-model, (B) using two separate analyses of pairs of populations, and (C) under a non-hierarchical F model performed on four populations, thus ignoring the underlying hierarchical population structure. The colors indicate the inferred model: convergent evolution (yellow), selection only in

the first group (blue), selection only in the second group (red), and no selection (black). Note that we use purple in the C panel, as this approach does not allow one to distinguish between different models of selection. For better visualization, we only plot 10,000 neutral loci among the 90,000 simulated, but the missing data show a very similar pattern.

Supporting information

Figure S1. Hierarchical F-model.

Directed acyclic graph describing the Bayesian formulation of the hierarchical F-model with 12 populations clustered in three groups at a given locus i . Square nodes represent data and circles represent model parameters to be estimated. Dashed circles represent population allele frequencies, which are analytically integrated using a Dirichlet-multinomial distribution (see method description). Lines between the nodes represent direct stochastic relationships within the model.

Figure S2. Original F-model.

Directed acyclic graph describing the Bayesian formulation of the original F-model with 12 populations at a given locus i . Square nodes represent data and circles represent model parameters to be estimated. Dashed circles represent population allele frequencies, which are analytically integrated using a Dirichlet-multinomial distribution (see method description). Lines between the nodes represent direct stochastic relationships within the model.

Table S1. Significant regions under altitudinal selection in Asian and Amerindian populations identified using the sliding windows approach.

The 25 genomic regions identified correspond to the vertical grey bars in Figure 3. We report the closest genes within 250kb from a significant SNP ($q\text{-value} < 0.01$) in each region. The corresponding SNPs for each gene are also reported. We highlight in bold the genes and regions discussed in the text and in Figure 3.

Table S2. Gene sets enriched for signals of convergent adaptation before pruning ($q\text{-value} < 0.05$).

We highlight in bold the only two gene sets that remain significant after the pruning procedure,

which consists in removing overlapping genes from less significant gene sets and retesting in an iterative manner.

Table S3. Results of the gene set enrichment approach for the "*Fatty Acid Omega Oxidation*" cluster.

We report the list of genes member of all gene sets in the "*Fatty Acid Omega Oxidation*" cluster (see Figure 4). The only remaining significant gene set after pruning ("*Fatty Acid Omega Oxidation*") is highlighted in yellow. For each gene we report the retained SNP (see Text S1), the distance to the gene, and the corresponding statistic used (*1-q-value*).

Text S1. Gene set enrichment method description,

Tables

Table 1: Result of the simulated data analyses.

SNP category	Selection parameters	Number of SNPs	New hierarchical model				Two separate analyses				Island model	
			Neutral	Selection			Neutral	Selection			Neutral	Selection
				Convergent	Group 1	Group 2		Convergent	Group 1	Group 2		
Convergent	$\alpha_i = 3$	2500	1824	306	148	222	1891	47	235	327	2127	373
	$\alpha_i = 5$	2500	550	1515	188	247	634	643	554	669	1048	1452
Group 1	$\alpha_{i1} = 3 \quad \alpha_{i2} = 0$	2500	2206	19	275	0	2214	0	285	1	2367	133
	$\alpha_{i1} = 5 \quad \alpha_{i2} = 0$	2500	1308	65	1127	0	1307	0	1192	1	1860	640
Neutral	$\alpha_{i1} = \alpha_{i2} = 0$	90,000	89971	0	4	25	89,970	0	4	26	88,861	1139
All		100,000	95859	1905	1742	494	96,016	690	2270	1024	96,263	3737

Figure 1.

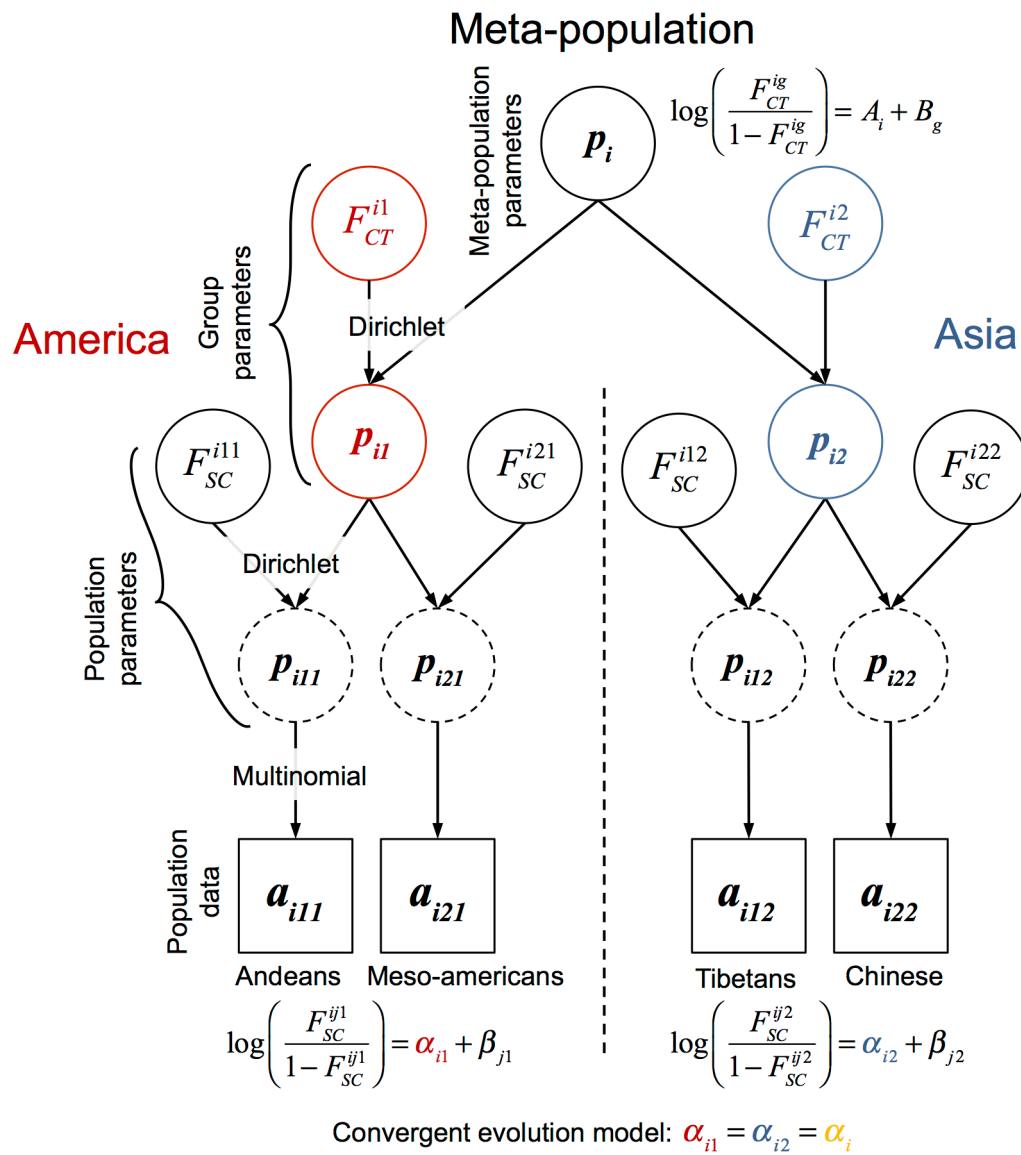


Figure 2.

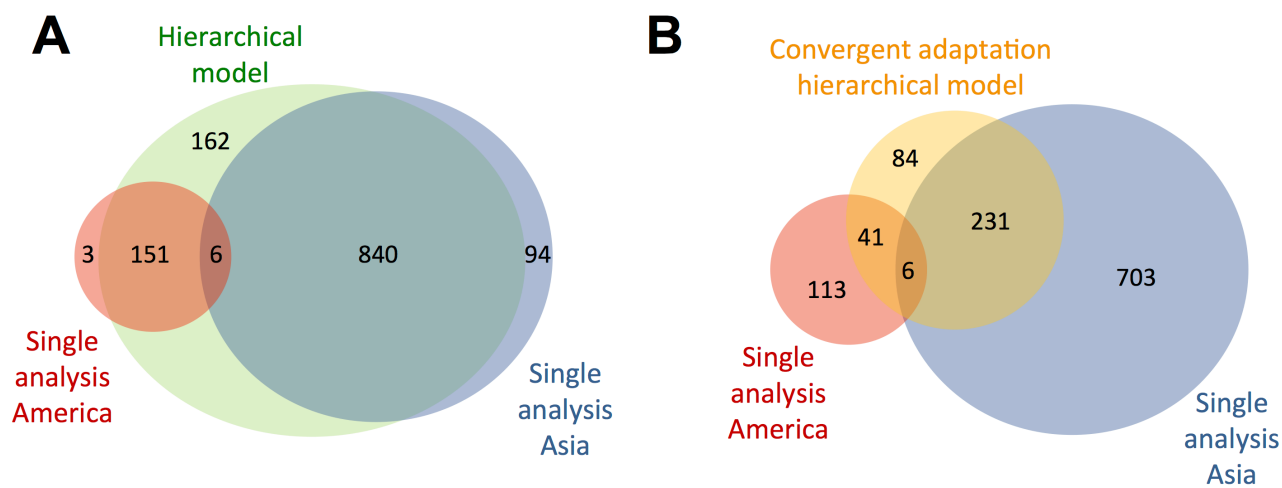


Figure 3.

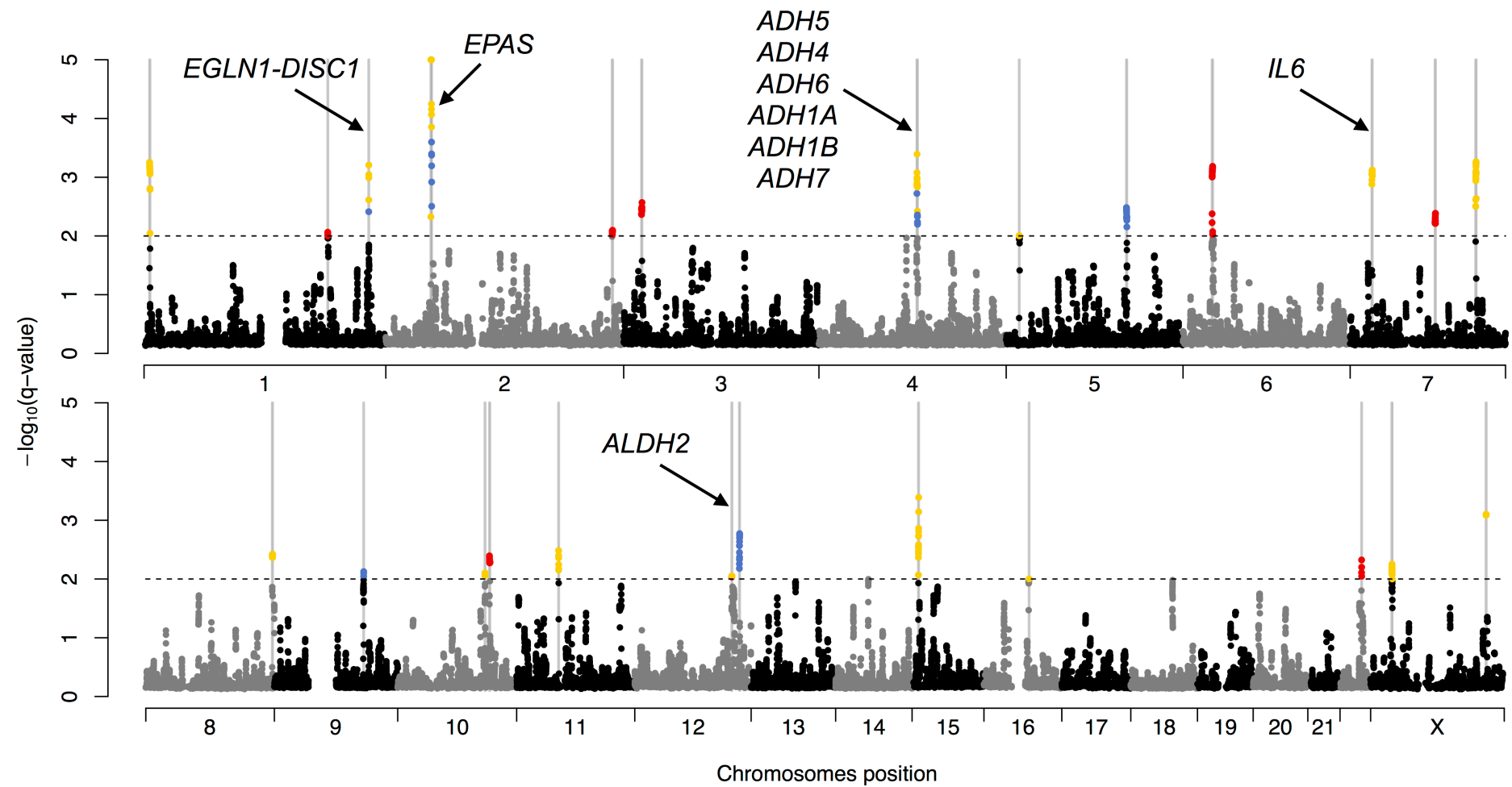


Figure 4.

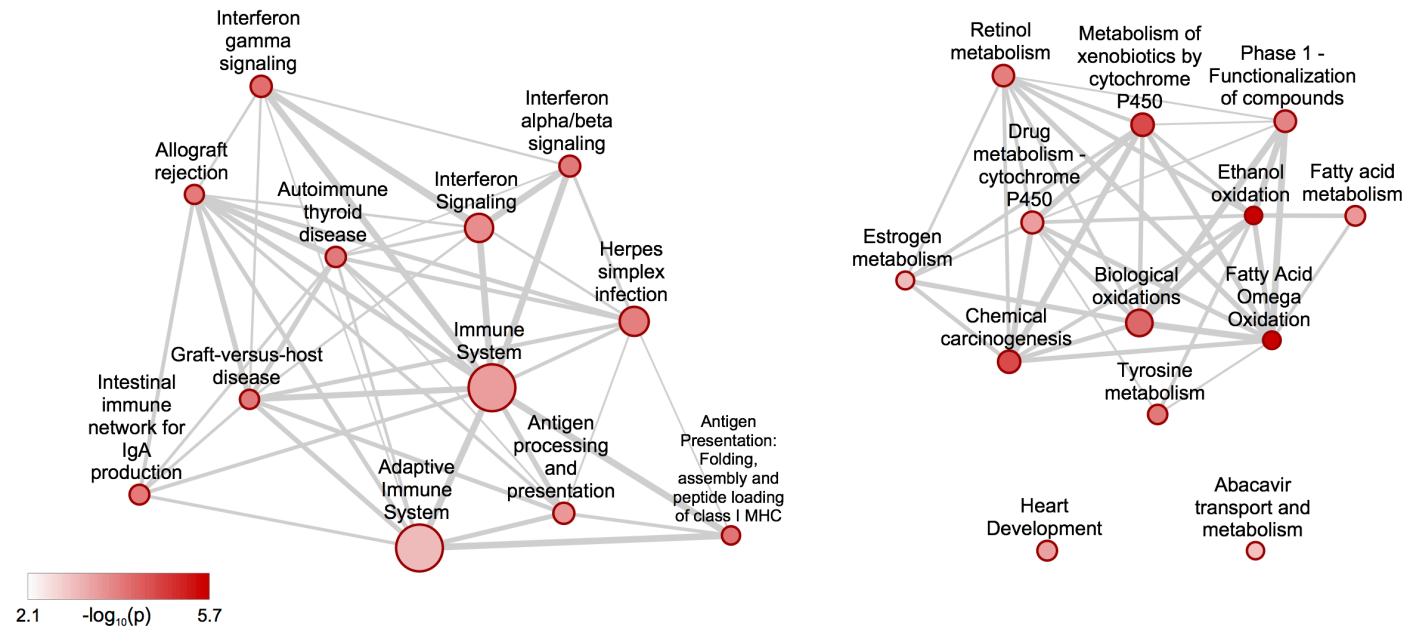


Figure 5.

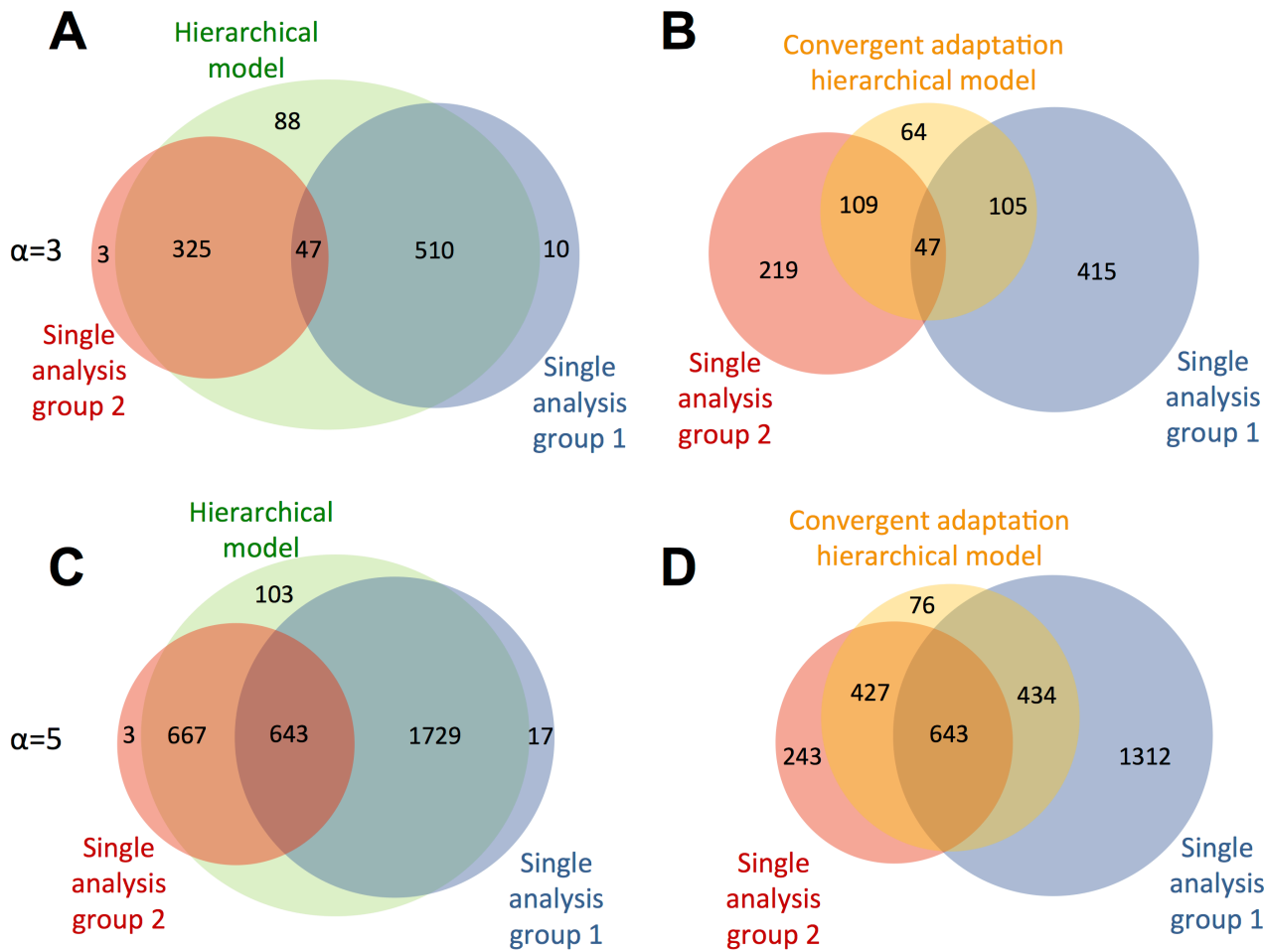


Figure 6.

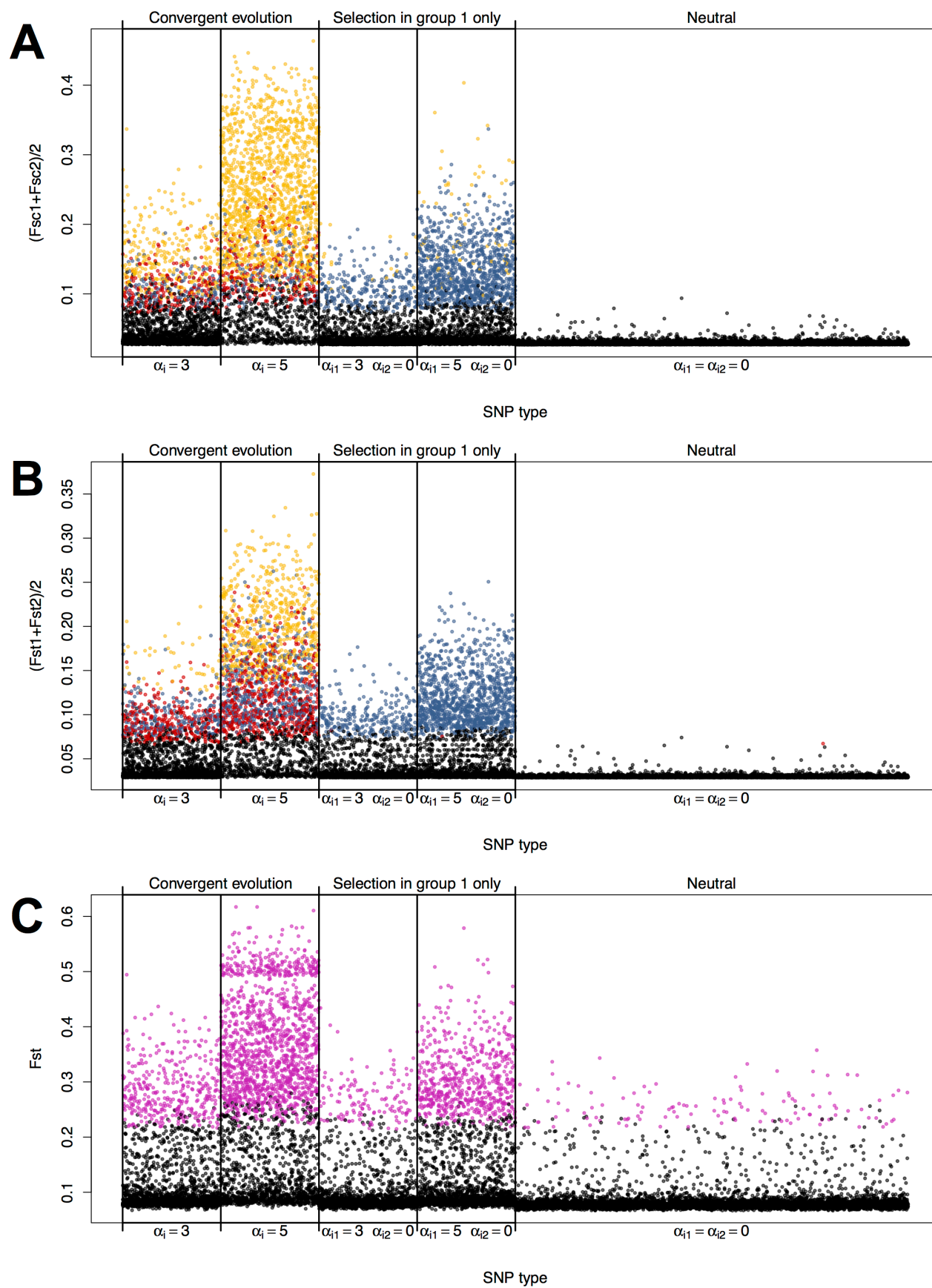


Figure S1.

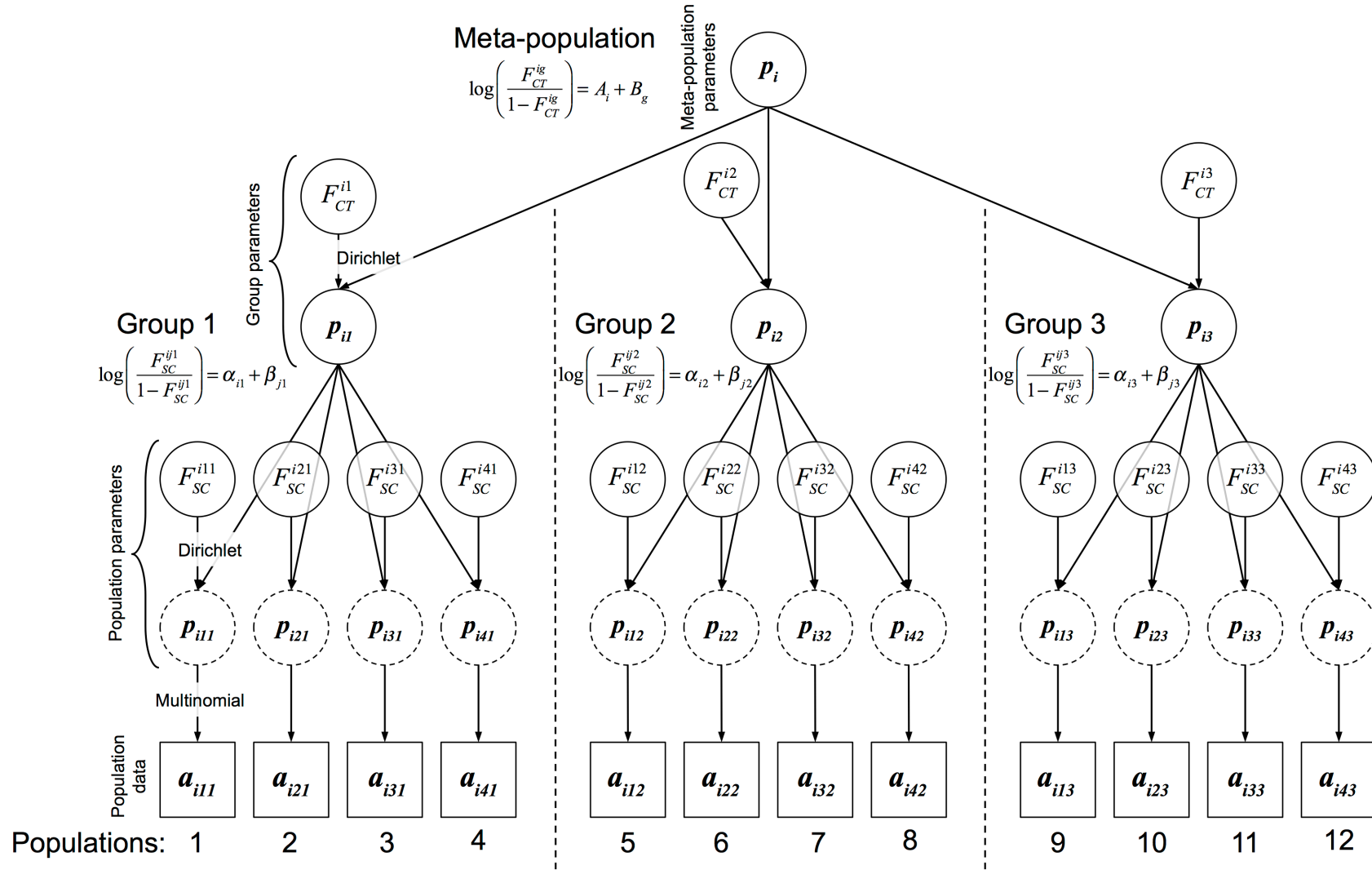


Figure S2.

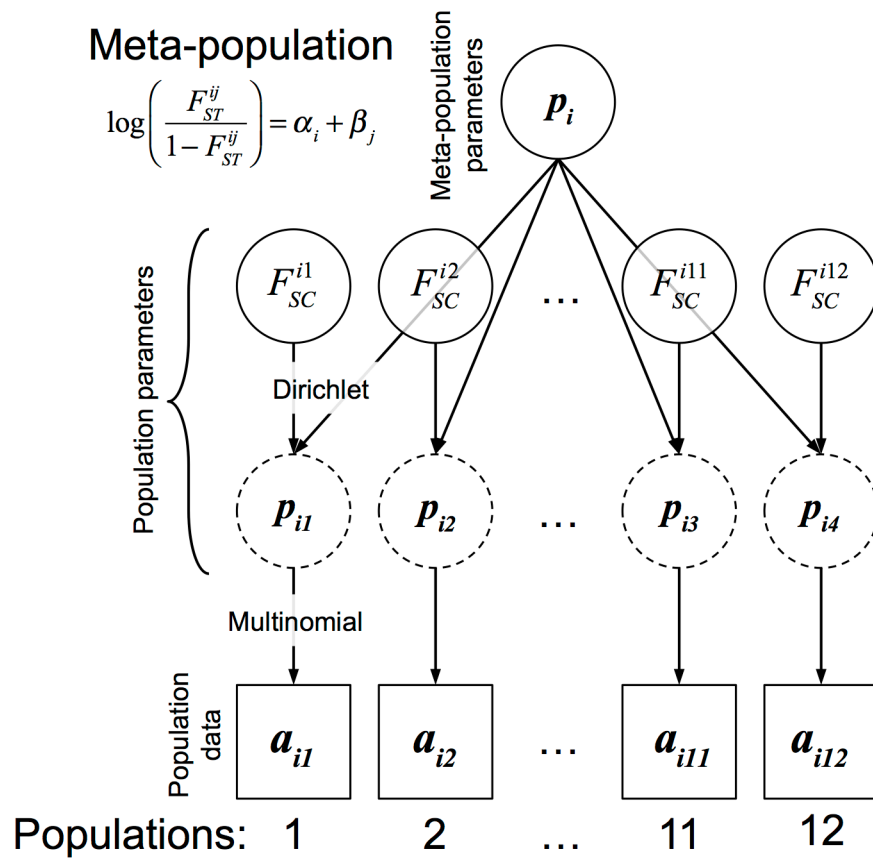


Table S1.

Chromosome	Region ID	Closest gene to significant SNPs (<250kb)	SNPs in region with qvalue<0.01
1	1	ZBTB48 KLHL21 THAP3 DNAJC11	rs731024 rs6682920 rs6678681 rs742394 rs200457 rs12137794 rs7549198 rs156985 rs3827729 rs399242 rs11122119 rs277671 rs277681
	2	-	-
	3	TRIM67 TSNAX-DISC1	rs12096847 rs12063614 rs12563076 rs11122250 rs6541261 rs16854592 rs12058372 rs12058117
2	4	PRKCE EPAS1 LOC388946 ATP6V1E2 RHOQ PIGF CRIPT SOC55	rs2594489 rs8179696 rs17034920 rs2121266 rs17034950 rs7571879 rs4952819 rs7582701 rs13419896 rs9679290 rs6758592 rs6715787 rs6544888 rs6756667 rs7589621 rs4953360 rs6755594 rs1374749 rs10178633 rs11675232 rs4953361 rs3088359 rs11678465 rs7583392 rs7594278 rs7571218 rs13006131 rs1868092 rs13424253 rs4450662 rs1109286 rs1447563 rs2346177 rs4953372 rs7587138 rs7556828 rs13003074 rs12622818 rs12619696 rs1542271 rs13032473 rs4953385 rs2276555 rs4953388 rs13001507 rs6741821 rs11125079 rs4953396 rs12470532 rs3814045 rs11676473 rs17035292 rs1901263 rs4953402 rs13000706 rs6735530 rs8390 rs3087822 rs12104572 rs12105006 rs7599097 rs10514802 rs10179861 rs10204096 rs4952838 rs12328738 rs6742593 rs991861 rs10209278 rs2346877 rs1319498 rs1378763 rs930853 rs10178013 rs10495936 rs880292 rs7596521 rs7584870 rs7562173 rs11695058
	5	SP140L	rs13012615 rs12694851 rs12694855 rs13383946 rs12694858 rs10184953 rs4364013 rs4577270
3	6	SATB1	rs336614 rs453585 rs4130090 rs11128893 rs13084808 rs7633180 rs6799759 rs12106877
4	7	TSPAN5	rs10031904 rs2178125 rs7685402
		EIF4E METAP1 ADH5 ADH4 ADH6 ADH1A ADH1B ADH7	rs1373244 rs7684429 rs17595102 rs1869458 rs2851275 rs3805322 rs2051428 rs6839510 rs1229966 rs3811802 rs1442488 rs969804 rs284793 rs284789 rs284787
5	8	DNAH5	rs2034221 rs1445691 rs2166337 rs1017573 rs2896110 rs30171 rs2652768 rs1354191
	9	CEP120 CSNK1G3	rs890933 rs11959808 rs7712330 rs7714655 rs10038345 rs12515732 rs11241705 rs10478591
6	10	ZFP57	rs3131847 rs3129045 rs2747442 rs3117294
		HLA-G	rs3115628

		HLA-A	rs2523969 rs2523957 rs2256919
		HLA-J	rs2735069
		RNF39	rs7382061
		TRIM31	rs6457144 rs9261394 rs4959041
		TRIM40	rs2517592 rs9261442 rs9261446 rs1541270 rs9261471 rs2857435 rs2857439 rs9261488 rs9261489 rs9261491 rs757262 rs757259
			rs1573298 rs9261518 rs9261519
		TRIM15	rs9261539
		TRIM26	rs1042338 rs3132671 rs2844775 rs3130391 rs3132666
		HLA-L	rs2844780
7	11	IL6	rs2961299 rs2905324 rs1006001 rs2961304 rs2961309 rs1548418 rs2961312 rs6946864 rs6969502 rs4719711 rs1404008
			rs6461662 rs6963591 rs1880242 rs2066992 rs2069852 rs7802277
		RPS26	rs4722175 rs9639435 rs9639436
	12	ABCB1	rs1045642 rs6949448 rs4148738 rs10808072 rs2235033 rs1202169 rs1202168 rs1202184
	13	OPN1SW	rs1868774
		IRF5	rs4728142 rs3807306
		TNPO3	rs12531711 rs12531054 rs17424602
		TPI1P2	rs12537496 rs13232316 rs17340542 rs13227095
8	14	PSCA	rs9297976
		LY6K	rs2164308 rs2082801 rs1469811 rs10956986
		GML	rs2717586 rs439747
9	15	TGFBR1	rs868
		SEC61B	rs894674 rs920771 rs7032399 rs7040144
10	16	ARHGAP19	rs793519
		RRP12	rs6584123
		MMS19	rs872106
		UBTD1	rs10882949 rs10882950 rs10882951
	17	NOLC1	rs7897
		C10orf26	rs2250580 rs2249845 rs7069489 rs10786708 rs7079231 rs549466 rs630185 rs2482496 rs2254093 rs2482506
11	18	RAPSN	rs17198158
		NDUFS3	rs4147730
		MTCH2	rs4752786
		AGBL2	rs4752791
		NUP160	rs2305982 rs6485788 rs4752797 rs7924699 rs1872167
12	19	CUX2	rs933307 rs1362006 rs7300860
		SH2B3	rs739496
		ATXN2	rs1029388
		ACAD10	rs11066019
		ALDH2	rs4767944
		MAPKAPK5	rs4346023
	20	CCDC64	rs11829349 rs11064983 rs7302874 rs12311327

15	21	OCA2 HERC2	rs4778210 rs1800414 rs12593141 rs2305252 rs3794602 rs3829488 rs16950821 rs895828 rs7179419 rs916977
16	22	TOX3 TPM3	rs3104767 rs3112625 rs12929797 rs3104780 rs3104784 rs3104800 rs3112609
22	23	TNRC6B MKL1	rs8138982 rs12485003 rs17001819 rs133054
X	24	SMEK3P	rs2218675 rs6418587 rs12859748 rs4240089 rs7061153 rs4297201 rs4357455 rs4829056
	25	PLAC1	rs5933443 rs5933446 rs5930658 rs5978032 rs5930660
		FAM122B	rs5933454 rs5933455 rs2355307 rs13440516

Table S2.

Rank	Set size	p-value	q-value	Set name
1	15	2.00E-06	0.001087	Fatty Acid Omega Oxidation
2	10	2.00E-06	0.001087	Ethanol oxidation
3	78	2.20E-05	0.006520	Metabolism of xenobiotics by cytochrome P450
4	76	2.40E-05	0.006520	Chemical carcinogenesis
5	131	6.20E-05	0.009160	Biological oxidations
6	62	7.60E-05	0.009160	Interferon gamma signaling
7	23	9.20E-05	0.009160	Antigen Presentation: Folding, assembly and peptide loading of class I MHC
8	49	1.00E-04	0.009160	Autoimmune thyroid disease
9	37	1.00E-04	0.009160	Tyrosine metabolism
10	35	1.02E-04	0.009160	Graft-versus-host disease
11	58	1.06E-04	0.009160	Interferon alpha/beta signaling
12	34	1.14E-04	0.009160	Allograft rejection
13	43	1.16E-04	0.009160	Intestinal immune network for IgA production
14	165	1.18E-04	0.009160	Herpes simplex infection
15	64	1.36E-04	0.009853	Retinol metabolism
16	66	1.46E-04	0.009916	Phase 1 - Functionalization of compounds
17	153	2.00E-04	0.012785	Interferon Signaling
18	41	2.88E-04	0.016953	Fatty acid metabolism
19	60	3.04E-04	0.016953	Antigen processing and presentation
20	945	3.12E-04	0.016953	Immune System
21	70	3.60E-04	0.018630	Drug metabolism - cytochrome P450
22	42	3.98E-04	0.019660	Heart Development
23	17	9.16E-04	0.043017	Estrogen metabolism
24	557	9.50E-04	0.043017	Adaptive Immune System
25	10	1.08E-03	0.047121	Abacavir transport and metabolism

	SULT2B1		19	rs2665582		0	0.28731
	GALM		2	rs12619390		0	0.28123
	GLYAT		11	rs472569		14252	0.27873
x	UGT2B7		4	rs12645107		10294	0.27335
	ACSL3		2	rs6756323		0	0.27255
	SULT1C4		2	rs17190797		12256	0.27007
	COMT		22	rs174699		0	0.2663
x	ADH1C		4	rs1614972		0	0.26505
	CYP11B2		8	rs4736354		11509	0.26504
	CYP24A1		20	rs17217119		27398	0.26455
	CBR3		21	rs2835288		6969	0.2632
	GCDH		19	rs11085825		0	0.25791
x	EPHX1		1	rs17571067		0	0.25651
x	MGST2		4	rs8192004		221	0.25594
	ALDH7A1		5	rs6862379		19248	0.25458
	PGAM1		10	rs11189214		5019	0.24525
	TBXAS1		7	rs7806848		0	0.2445
	CPT1A		11	rs3019607		0	0.24162
	BPGM		7	rs10488453		41034	0.24073
	MIF		22	rs4820571		5564	0.23939
	CYP2R1		11	rs1553006		9199	0.23262
	RGN	X		rs6521178		0	0.23234
	ACADM		1	rs1463812		8106	0.23023
	ACAT2		6	rs927450		893	0.21685
	CYP2J2		1	rs11572223		0	0.21628
	PGM2		4	rs6531588		0	0.215
	DDC		7	rs7803788		0	0.21412
x	GSTP1		11	rs7952081		26133	0.2121
x	UGT2B11		4	rs4438816		34377	0.20968
	FAHD1		16	rs3848346		0	0.209
	CYP4A22		1	rs12564103		8172	0.20756
	ADPGK		15	rs11631352		10347	0.20702
	STS	X		rs7884548		45611	0.20644
x	GSTA2		6	rs4715309		28632	0.19774
	MAOB	X		rs5952352		0	0.19713
	PNPLA4	X		rs7888492		6078	0.19243
x	UGT2B10		4	rs1458232		3868	0.18923
	FMO3		1	rs9970392		1699	0.18754
x	CYP1B1		2	rs162556		3131	0.18684
	CYP7B1		8	rs6472155		18859	0.18608
	ALDH3A2		17	rs4646798		0	0.18088
	GGT7		20	rs17092148		0	0.18042
	CYP4F11		19	rs2072269		0	0.18022
x	AKR1C2		10	rs7915338		2307	0.17941
	CYP4F8		19	rs3764563		1826	0.17572
	IL4I1		19	rs1290743		0	0.17491
x	CYP2B6		19	rs16974799		0	0.17471
x	MGST1		12	rs10846364		49997	0.17314
	GGCT		7	rs38433		9205	0.17274
	AOX1		2	rs7563911		40066	0.17263
	G6PC		17	rs161634		0	0.17207
	SULT1B1		4	rs1847366		0	0.17034
	CYP4F12		19	rs6311193		0	0.16868
	ALDOA		16	rs11642740		3756	0.16786
	MTR		1	rs4006372		10534	0.16752
x	UGT2B4		4	rs13151633		33826	0.166
x	UGT2B28		4	rs13140179		10512	0.1656
x	CCBL1		9	rs12551834		0	0.16554
	DLD		7	rs6965674		3484	0.16398
x	CYP3A43		7	rs1403195		1305	0.16398
	TPO		2	rs17091737		0	0.16221
	MAT2B		5	rs297954		40708	0.16187
	CYP19A1		15	rs6493496		0	0.16155

		ACSL4	X	rs5943427	0	0.15148
		NNMT		11 rs4938091	23905	0.15101
x		CCBL2		1 rs3753683	0	0.15083
		ACSBG1		15 rs1105401	0	0.14993
		DLAT		11 rs7949136	0	0.14944
x		ALDH3B1		11 rs581105	0	0.14814
		ACSM1		16 rs433598	0	0.14765
		SULT4A1		22 rs138099	0	0.14752
		CYP11B1		8 rs7818826	0	0.1457
x		GSTA3		6 rs534232	7064	0.14447
		NAT1		8 rs10503610	19557	0.14389
		ACSS1		20 rs761720	0	0.14386
		GAPDH		12 rs7971637	256	0.14324
x		CYP4A11		1 rs1502931	25714	0.14251
		UGT2A1		4 rs6600789	43023	0.14213
		POMC		2 rs12619264	30244	0.1408
		EHHADH		3 rs16859799	1700	0.14073
		ACSL5		10 rs1887139	13477	0.14004
x		CYP2C8		10 rs11188183	33032	0.14004
		CYP26B1		2 rs12477733	9326	0.13939
		ALDH1A2		15 rs16977865	1211	0.13932
		ACAA1		3 rs2239621	0	0.13844
		ACAA2		18 rs7239952	29150	0.13761
		AKR1C4		10 rs12247748	0	0.13696
		PAPSS2		10 rs3824721	0	0.13696
		ALDOB		9 rs7467699	1985	0.13695
		TPMT		6 rs2842934	0	0.13692
x		ALDH1A3		15 rs4965339	33158	0.1368
		G6PC2		2 rs16856159	0	0.13671
x		HSD11B1		1 rs2298930	0	0.13665
		FMO1		1 rs7551531	2783	0.13665
		TYR		11 rs10830253	0	0.13443
		PTGS1		9 rs10306150	0	0.1329
		ACADS		12 rs509152	8738	0.12971
		HPD		12 rs11043218	0	0.12971
x		CYP3A7		7 rs2687145	0	0.12842
		CYP2C9		10 rs10509679	0	0.1278
		SDR16C5		8 rs4075155	17953	0.1271
		ACADSB		10 rs10794583	0	0.12652
		FMO2		1 rs7542361	0	0.1262
		RDH8		19 rs4322768	23944	0.12602
		TAT		16 rs4788811	14211	0.12535
		PGM1		1 rs10889433	0	0.12495
x		MGST3		1 rs16847570	26812	0.12495
		CYP4B1		1 rs12082811	38784	0.12495
x		SULT2A1		19 rs2932766	770	0.12483
		HK2		2 rs6732614	0	0.12437
		TYRP1		9 rs16929332	21693	0.1239
		RDH12		14 rs12882315	0	0.12343
x		ALDH3B2		11 rs4930497	26916	0.1219
		CYP26A1		10 rs7905939	3466	0.12182
x		GSTO2		10 rs157081	0	0.12182
		GSTA1		6 rs9395826	11026	0.12153
		GCK		7 rs741038	0	0.12121
		RPE65		1 rs7515572	42939	0.12117
x		PTGS2		1 rs16825773	19572	0.12117
		AWAT2	X	rs1409995	0	0.12044
x		UGT2A2		4 rs17147521	0	0.12018
		SLC35D1		1 rs2815378	0	0.12005
		PTGIS		20 rs495146	0	0.11976
		LDHA		11 rs11024636	10445	0.11939
		MAT2A		2 rs6705971	4684	0.11929
		CYP4F3		19 rs4807964	2880	0.11904
		RDH10		8 rs6472767	32248	0.11877
		ALDH9A1		1 rs10918241	0	0.11766
		LRAT		4 rs156500	3045	0.11752
		ACOX1		17 rs7217955	0	0.11751
		PCK1		20 rs2179706	0	0.11741
		DHRS4		14 rs3742492	0	0.11705
x		GSTO1		10 rs11191966	10840	0.11699

x									GPI	19	rs8191425	0	0.0897
	x	x	x	x	x	x	x	x	CYP3A4	7	rs6945984	6255	0.08945
	x								RDH5	12	rs3138139	0	0.08866
		x	x					x	GSTT2	22	rs140245	3605	0.0885
			x						FMO5	1	rs10900326	0	0.08692
		x	x	x				x	GSTM4	1	rs542338	5932	0.08627
				x					AHCY	20	rs6088466	13926	0.08622
x									HADH	4	rs221347	0	0.0856
		x	x					x	GSTT2B	22	rs2858908	3825	0.08455
		x	x	x				x	GSTM5	1	rs11807	0	0.08403
x									LDHAL6A	11	rs11024671	0	0.08301
	x								ACAT1	11	rs10890817	0	0.08278
x									LDHAL6B	15	rs3816814	0	0.08271
				x	x				CYP27B1	12	rs4646536	0	0.08271
	x								METTL2B	7	rs4731470	15932	0.08121
		x		x	x				CYP2S1	19	rs1645684	31842	0.08088
		x	x					x	GSTK1	7	rs10248147	19499	0.08083
				x	x				CYP2U1	4	rs3756271	0	0.08076
	x								DHRS4L2	14	rs1811890	0	0.08065
	x								AOC2	17	rs16968038	46	0.08034
		x	x	x			x	x	GSTM1	1	rs2071487	0	0.07994
x									PDHB	3	rs7231	0	0.07855

Text S1.

Gene set enrichment analysis method

To find signals of selection at the pathway level we applied a gene set enrichment approach as described by Daub et al. (2013). This method tests whether the genes in a gene set show a shift in the distribution of a selection score. In our case we take as selection score $s_{conv} = 1 - q_{conv}$, where q_{conv} is the q-value of a SNP computed from the probability of convergent selection. For the enrichment test we need one s_{conv} value per gene, we therefore transformed the SNP based scores to gene based scores. We first downloaded 19,683 protein coding human genes, located on the autosomes and on the X chromosome, from the NCBI Entrez Gene website (Maglott et al. 2011, <http://www.ncbi.nlm.nih.gov/gene>, downloaded on February 7, 2013). Next we converted the SNPs to hg19 coordinates. 670 SNPs could not be mapped, resulting in 631,674 remaining SNPs. These SNPs were assigned to genes: if a SNP was located within a gene transcript, it was assigned to that gene; otherwise it was assigned to the closest gene within 50kb distance. For each gene, we selected the highest s_{conv} value of all SNPs assigned to this gene. After removing 2,411 genes with no SNPs assigned, a list of 17,272 genes remained.

We downloaded 2,402 gene sets from the NCBI Biosystems database (Geer et al. 2010, <http://www.ncbi.nlm.nih.gov/biosystems>, downloaded 7 Feb 2013). After discarding genes that were not part of the aforementioned gene list, removing gene sets with less than 10 genes and pooling (nearly) identical gene sets, 1,339 sets remained that served as input in our enrichment tests.

We computed the SUMSTAT (Tintle et al. 2009) score for each set, which is the sum of the s_{conv} values of all genes in a gene set. Gene sets with a high SUMSTAT score are likely candidates for convergent selection. To assess the significance of each tested gene set, we compared its SUMSTAT score with a null distribution of SUMSTAT scores from random gene sets (N=500,000) of the same size. We could not approximate the null distribution with a normal distribution (as applied in Daub et al. 2013), as random gene sets of small to moderate size produced a skewed SUMSTAT distribution. Taking the highest s_{conv} score among SNPs near a gene can induce a bias, since genes with many SNPs are more likely to have an extreme value assigned. To correct for this possible bias we placed each gene in a bin containing all genes with approximately the same number of SNPs and constructed the random gene sets in the null distribution in such a way that they were composed of the same number of genes from each bin as the gene set being tested. To

remove overlap among the candidate gene sets, we applied a pruning method where we assign iteratively overlapping genes to the highest scoring gene set. As these tests are not independent anymore, we empirically estimated the q-value of these pruned sets. All sets that scored a q-value <5% (before and after pruning) were reported.

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